

Testing network correlation efficiently via counting trees

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Joint work with
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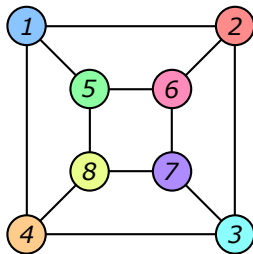
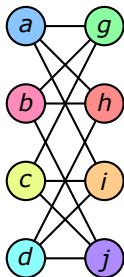
Given two graphs A and B , decide whether $A \cong B$, i.e., there exists a bijection $\pi : V(A) \rightarrow V(B)$ such that

$$(u, v) \in E(A) \Leftrightarrow (\pi(u), \pi(v)) \in E(B)$$

Graph isomorphism

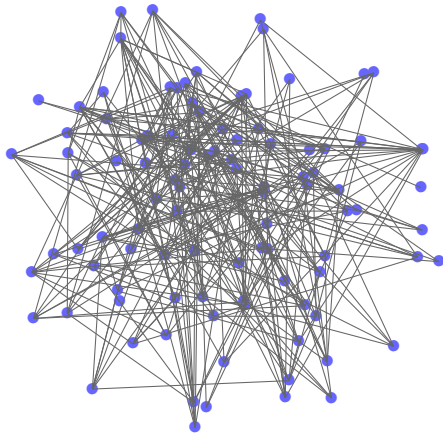
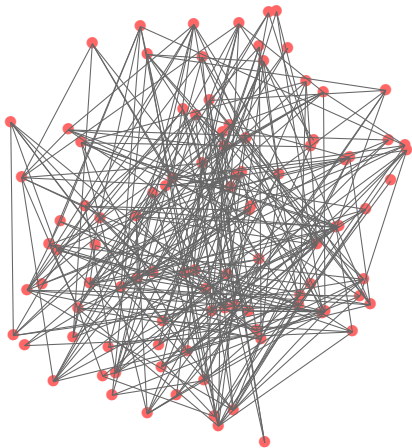
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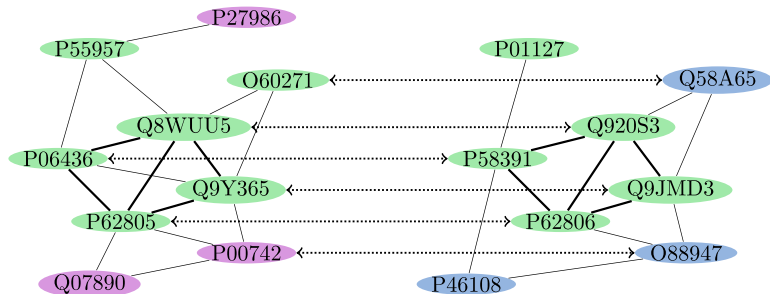
- Not known to be solvable in polynomial time in worst case

Two key challenges



- **Statistical:** two graphs may be correlated but not exactly isomorphic
- **Computational:** # of possible node mappings is $n!$ ($100! \approx 10^{158}$)

Motivation: Protein-protein interaction network



Human network

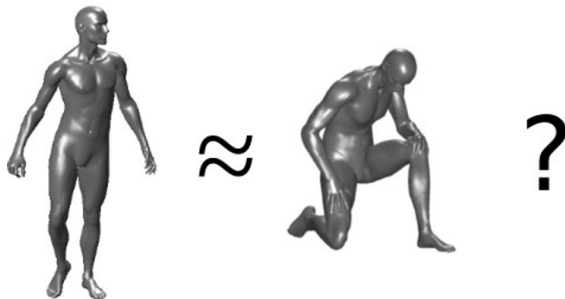
Mouse network

[Kazemi-Hassani-Grossglauer-Modarres '16]

Ontology: Assess the correlation of two biological networks in two different species based on network topology

Motivation: Deformable shape matching

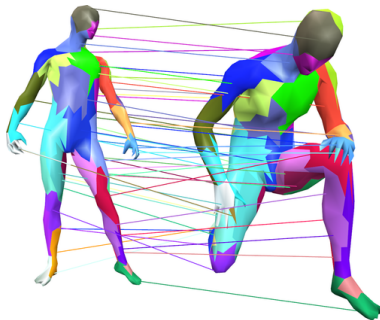
A fundamental problem in computer vision: Detect similar objects that undergo different deformations



Shape REtrieval Contest (SHREC) dataset

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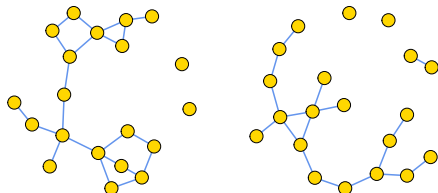
3-D shapes \rightarrow geometric graphs (features \rightarrow nodes, distances \rightarrow edges)
Determine whether two graphs are similar in topologies

Beyond worst-case: noisy graph isomorphism test

Definition

\mathbb{H}_0 : A and B are independent Erdős-Rényi graphs $\mathcal{G}(n, q)$

\mathbb{H}_1 : A and $B^\pi = (B_{\pi(i)\pi(j)})$ are correlated $\mathcal{G}(n, q)$ with correlation parameter ρ , conditional on a uniform permutation π

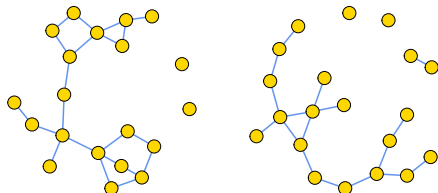


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- Hypothesis testing aspect of **graph matching** (recover π under \mathbb{H}_1)
- The inherent edge correlation is obscured by latent node mapping π
- The test needs to rely on **graph invariants** (e.g. # of edges or triangles)

Key questions

Goal: Design a test statistic $f(A, B)$ with threshold $\tau \in \mathbb{R}$ to achieve consistent detection:

$$\lim_{n \rightarrow \infty} [\mathbb{P}_{\mathbb{H}_1} (f(A, B) < \tau) + \mathbb{P}_{\mathbb{H}_0} (f(A, B) \geq \tau)] = 0$$

Two key questions

- What is the statistical detection limit?
- Can we attain the limit in polynomial-time?

Statistical detection limit

Consider an equivalent model: under \mathbb{H}_1 , A and B are independently subsampled from a common parent graph $\mathcal{G}(n, p)$ with probability s :
 $q = ps$ and $\rho = \frac{s(1-p)}{1-ps}$

Theorem (Wu-X.-Yu '20)

Consistent detection is possible if and only if

$$s^2 \geq \begin{cases} \frac{2 \log n}{np \left(\log \frac{1}{p} - 1 + p \right)} & \text{if } p = n^{-o(1)} \text{ (dense)} \\ \Theta \left(\frac{1}{np} \right) & \text{if } p = n^{-\Omega(1)} \text{ (sparse)} \end{cases}$$

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- [\[Ding-Hu' 22\]](#) determined the sharp constant in the sparse graph case
- Achieved by computationally intractable statistic:

$$\text{MCS}(A, B) \triangleq \max_{\pi \in \mathcal{S}_n} \sum_{i < j} A_{ij} B_{\pi(i)\pi(j)} \quad (\text{max common subgraph})$$

- Question: How to test efficiently?

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- Question: How to test efficiently? **subgraph count**

Testing via subgraph counts

Let $X_H(G)$ denote the number of copies of H in G

- Under \mathbb{H}_0 : $X_H(A)$ and $X_H(B)$ are **independent**
- Under \mathbb{H}_1 : $X_H(A)$ and $X_H(B)$ are **correlated**

Implication: Test based on the correlation between $X_H(A)$ and $X_H(B)$

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First attempt: H is a single edge [Wu-X.-Yu '20]

$$\frac{X_H(A) - X_H(B)}{\sqrt{n(n-1)q(1-q)}} \approx \begin{cases} \mathcal{N}(0, 1) & \text{Under } \mathbb{H}_0 \\ \mathcal{N}(0, 1 - \rho) & \text{Under } \mathbb{H}_1 \end{cases}$$

\Rightarrow consistent detection demands $\rho \rightarrow 1$

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Lesson: need to count more “distinguishing” subgraph

Counting a flock of black swans

Counting a family \mathcal{H} of **rare** subgraphs [Barak-Chou-Lei-Schramm-Sheng'19]

- H is rare: $\mathbb{E}[X_H(A)] = O(1) \Rightarrow$ unlikely to co-occur under \mathbb{H}_0
- \mathcal{H} needs to be “rich” to ensure co-occurrence under \mathbb{H}_1
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Construction of \mathcal{H} based on **random d -regular graphs** of size $K = \rho^{-C}$
 \Rightarrow correctly tell \mathbb{H}_0 vs. \mathbb{H}_1 w.p. 0.9 in time n^K , if

$$\rho = \Omega(1) \quad \text{and} \quad nq \in \left[n^\epsilon, n^{1/153} \right] \cup \left[n^{2/3}, n^{1-\epsilon} \right]$$

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Limitations:

- nq needs to fall into specific range and does not cover sparse graphs
- The test error is not vanishing
- The runtime grows polynomially at rate of $\rho^{-C} \Rightarrow$ too expensive for large networks

Our idea: counting large trees

Our strategy: count a family \mathcal{T} of trees of K edges with K growing in n

e.g. $K = 4$, $\mathcal{T} = \left\{ \begin{array}{c} \circ - \circ - \circ - \circ - \circ, \\ \begin{array}{c} \circ \\ | \\ \circ - \circ - \circ \\ | \\ \circ \end{array}, \\ \begin{array}{c} \circ \\ | \\ \circ - \circ \\ | \\ \circ \end{array} \end{array} \right\}$

- # of vertices = # of edges + 1 \Rightarrow abundant even in sparse graphs
- \mathcal{T} is “rich”: $|\mathcal{T}|$ grows exponentially in K
- Tree structure allows for efficient counting even when K grows in n

An obstacle: subgraph counts for different trees are still too correlated, especially when $nq \geq 1$

Our solution: counting **signed** trees

- Centered adjacency matrices: $\bar{A} = A - \mathbb{E}[A]$ and $\bar{B} = B - \mathbb{E}[B]$

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$$f_{\mathcal{T}}(A, B) \triangleq \sum_{H \in \mathcal{T}} \text{aut}(H) W_H(\bar{A}) W_H(\bar{B}),$$

where $\text{aut}(H)$ is the number of automorphism of H , and

$$W_H(M) \triangleq \sum_{S \cong H} \prod_{(i,j) \in S} M_{ij} \quad (\text{signed subgraph count})$$

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- Crucial orthogonal property:

$$\mathbb{E} [W_H(\bar{A}) W_{H'}(\bar{A})] = 0 \iff H \not\cong H'$$

Theorem (Mao-Wu-X.-Yu '21)

Suppose

$$nq \geq n^{-o(1)}, \quad \rho^2 > \alpha, \quad \omega(1) \leq K = O\left(\frac{\log n}{\log \log n}\right).$$

Then the testing error satisfies

$$\mathbb{P}_{\mathbb{H}_1}(f_{\mathcal{T}}(A, B) \leq \tau) + \mathbb{P}_{\mathbb{H}_0}(f_{\mathcal{T}}(A, B) \geq \tau) = o(1),$$

where $\tau = C\mathbb{E}_{\mathbb{H}_1}[f_{\mathcal{T}}(A, B)]$ for any fixed constant $0 < C < 1$.

- $\alpha \approx 0.33833$ is **Otter's constant** given by $|\mathcal{T}| \approx (1/\alpha)^K$
- $nq \geq n^{-o(1)}$ applies to very sparse regime and is necessary for existence of trees with $K = \omega(1)$ edges

$$f_{\mathcal{T}}(A, B) \triangleq \sum_{H \in \mathcal{T}} \text{aut}(H) W_H(\bar{A}) W_H(\bar{B}), \quad W_H(M) \triangleq \sum_{S \cong H} \prod_{(i,j) \in S} M_{ij}$$

- $\mathbb{E}_{\mathbb{H}_0}[f_{\mathcal{T}}(A, B)] = 0$ and by the orthogonal property,

$$\text{(SNR)} \quad \frac{(\mathbb{E}_{\mathbb{H}_1}[f_{\mathcal{T}}(A, B)])^2}{\text{Var}_{\mathbb{H}_0}[f_{\mathcal{T}}(A, B)]} = \rho^{2K} |\mathcal{T}| \approx \left(\frac{\rho^2}{\alpha}\right)^K$$

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- To drive the SNR to ∞ , need $\rho^2 > \alpha \approx 0.33833$
- Further show

$$\frac{(\mathbb{E}_{\mathbb{H}_1}[f_{\mathcal{T}}(A, B)])^2}{\text{Var}_{\mathbb{H}_1}[f_{\mathcal{T}}(A, B)]} \rightarrow \infty$$

Needs delicate analysis of correlations between subgraph counts

Computational guarantees

How to compute $W_H(M) \triangleq \sum_{S \cong H} \prod_{(i,j) \in S} M_{ij}$ efficiently?

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Efficient tree counts via color coding

How to compute $W_H(M) \triangleq \sum_{S \cong H} \prod_{(i,j) \in S} M_{ij}$ efficiently?

- Exhaustive search takes n^K time: super poly-time when $K \rightarrow \infty$
- Our solution: color coding [Alon-Yuster-Zwick '95]
 - ① Assign random color μ to each vertex from color set $[K+1]$ uniformly
 - ② Define *colorful* subgraph count

$$Y_H(M, \mu) \triangleq \sum_{S \cong H} \mathbf{1}_{\{S \text{ has } K+1 \text{ distinct colors}\}} \prod_{(i,j) \in S} M_{ij}$$

Key implication: $\mathbb{E}_\mu [Y_H(M, \mu)] = rW_H(M)$, where $r = \frac{(K+1)!}{(K+1)^{K+1}}$

- ③ Generate $1/r$ independent random colorings μ_t so that

$$W_H(M) \approx \sum_{t=1}^{1/r} Y_H(M, \mu_t) \triangleq \widehat{W}_H(M)$$

$Y_H(M, \mu)$ can be computed via dynamic programming in $n^2 e^{O(K)}$

Define an approximate statistic:

$$\tilde{f}_{\mathcal{T}} \triangleq \sum_{H \in \mathcal{T}} \text{aut}(H) \widehat{W}_H(\bar{A}) \widehat{W}_H(\bar{B})$$

Theorem (Mao-Wu-X.-Yu '21)

Under both \mathbb{H}_0 and \mathbb{H}_1 ,

$$\frac{\tilde{f}_{\mathcal{T}} - f_{\mathcal{T}}}{\mathbb{E}_{\mathbb{H}_1}[f_{\mathcal{T}}]} \xrightarrow{L_2} 0.$$

Moreover, $\tilde{f}_{\mathcal{T}}$ can be computed in time $n^{2+o(1)}$.

Implication: $\tilde{f}_{\mathcal{T}}$ satisfies the same statistical guarantee as $f_{\mathcal{T}}$

Statistical-computational gap

- Statistical limit: $\rho \approx \frac{1}{nq}$
- Tree counting limit: $\rho^2 > \alpha \approx 0.33833$
- There is a large computational gap, especially when the graphs are dense

Key question: Can we close the computational gap?

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Key question: Can we close the computational gap?

No, based on the low-degree hardness conjecture [Hopkins '18, Kunisky-Wein-Bandeira '19]: computationally hard when $\rho \leq \frac{1}{\text{polylog}(n)}$

Signed tree counting as a low-degree projection

- Space of functions on the hypercube $\{0, 1\}^{2\binom{n}{2}}$ endowed with

$$\langle f, g \rangle \triangleq \mathbb{E}_{\mathbb{H}_0} [f(A, B)g(A, B)]$$

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- Orthonormal polynomial basis [O'Donnell '4]

$$\phi_S(A, B) \triangleq \sigma^{-|S_1|-|S_2|} \prod_{(i,j) \in S_1} \bar{A}_{ij} \prod_{(k,\ell) \in S_2} \bar{B}_{k\ell}, \quad S \triangleq (S_1, S_2)$$

where $\sigma^2 = q(1 - q)$

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- Low-degree projection of likelihood ratio $L(A, B) = \frac{\mathbb{P}_{\mathbb{H}_1}(A, B)}{\mathbb{P}_{\mathbb{H}_0}(A, B)}$

$$f_{\mathcal{H}}(A, B) \triangleq \sum_{H \in \mathcal{H}} \sum_{S_1: S_1 \cong H} \sum_{S_2: S_2 \cong H} \langle L, \phi_S \rangle \phi_S$$

for given collection \mathcal{H} of unlabeled graphs

- “Optimal” degree- $2K$ polynomial of (A, B) :

$$f^* = \arg \max_{f: \text{degree}(f) \leq 2K} \frac{(\mathbb{E}_{\mathbb{H}_1} [f])^2}{\mathbb{E}_{\mathbb{H}_0} [f^2]}$$

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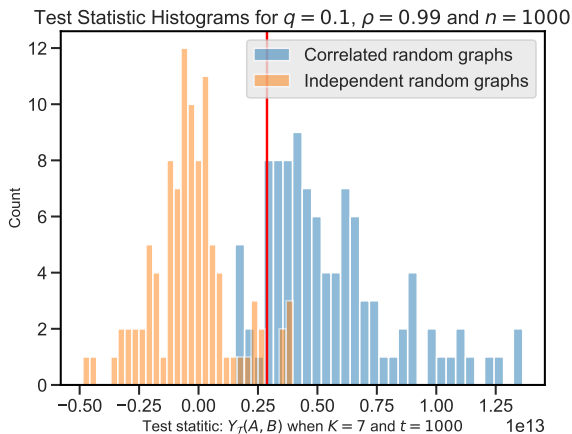
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- Low-degree hardness conjecture [Hopkins '18, Kunisky-Wein-Bandeira '19]:
All degree- $\log n$ polynomials fail when $\rho \leq \frac{1}{\text{polylog}(n)}$

Simulation study

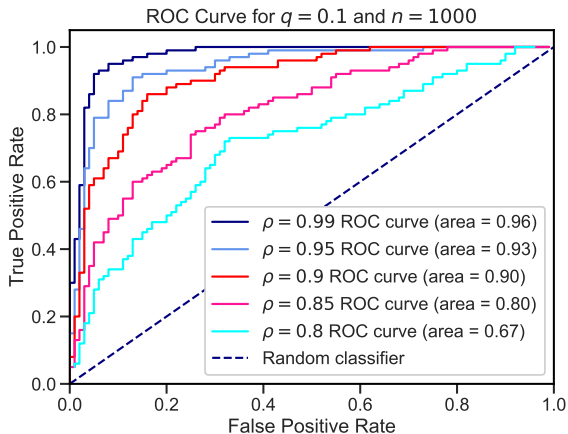
Simulation results

$n = 1000$, $K = 7$, $\rho = 0.99$



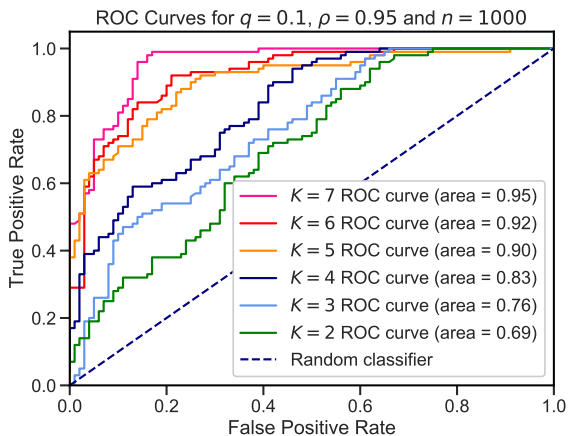
Type-I and type-II error are around 5% and 9%

$n = 1000$ and $K = 6$



Consistent with theoretical prediction: need $\rho \gtrsim |T|^{-\frac{1}{2K}} \approx 0.82$ for SNR exceeding 1, as $|\mathcal{T}| = 11$ when $K = 6$.

$n = 1000, q = 0.1, \rho = 0.95$



Performance of tree counting statistic improves as tree size K increases

Concluding remarks

- Efficient test via counting signed trees up to Otter's constant
- Low-degree hardness conjecture: no efficient test when $\rho \leq \frac{1}{\text{polylog}(n)}$

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Future work:

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- Recovery of node correspondence via signed tree signatures
- Testing network correlation in the presence of degree heterogeneity

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Reference

- C. Mao, Y. Wu, J. X., & S. H. Yu *Testing network correlation efficiently via counting trees*. [arXiv:2110.11816](https://arxiv.org/abs/2110.11816)
- Y. Wu, J. X., & S. H. Yu *Testing correlation of unlabeled random graphs*. [arXiv:2008.10097](https://arxiv.org/abs/2008.10097). To appear in *Annals of Applied Probability*